MAXIMUM LIKELIHOOD ESTIMATION OF THE COVARIANCES OF THE VECTOR MOVING AVERAGE MODELS IN THE TIME AND FREQUENCY DOMAINS

BY

F. AHRABI

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DEPARTMENT OF STATISTICS
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Maximum Likelihood Estimation of the Covariances of the Vector Moving Average Models in the Time and Frequency Domains

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Abstract

The vector moving average process is a stationary stochastic process \( \{y_t\} \) satisfying \( y_t = \sum_{i=0}^{q} A_i \varepsilon_{t-i} \), where the unobservable process \( \{\varepsilon_t\} \) consists of independently identically distributed random variables. The matrix parameters \( \Sigma^{(s)} = \sum_{t-s}^{\infty} y_t y'_{t+s} \), \( s = 0, 1, \ldots, q \) are estimated from the observations \( y_1, \ldots, y_T \). The likelihood function is derived under normality and to solve the maximum likelihood equations the Newton-Raphson and Scoring methods are used. The estimation problem is considered in the time and frequency domains. Asymptotic efficiency of the estimates is established.

Key words: Maximum likelihood estimation, vector moving average models, Newton-Raphson and Scoring iterative procedures, Time and Frequency Domains.
Maximum Likelihood Estimation of the Covariances of the Vector Moving Average Models in the Time and Frequency Domains

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1. Introduction

There have been a number of papers dealing with the estimation of the vector autoregressive moving average (VARMA) models. Hannan (1969, 1970) has considered the problem in the pure moving average case in the frequency domain. Nicholls (1976) has extended this to the case of VARMA models which include exogenous variables. Reinsel (1976) has considered the problem in the time domain. There have been other papers in this area, among them Akaike (1973), Tunnicliffe Wilson (1973), Kashyap (1970), Whittle (1963), and Osborn (1977). In all these papers the parameters of interest are the matrix coefficients of the vectors of observable and unobservable random variables and the common variance covariance matrix of the unobservable random variables.

This paper is concerned with the estimation of vector moving average models, but following Anderson (1975), Parzen (1971), and Clevenson (1970) we take as our parameters the autocovariance matrices of the observable random variables.

There is one paper by Newton (1975) which is primarily concerned with the usual parameterization of the VARMA models. However he derives estimates for the autocovariance matrices in the pure moving average case. His method is to regress the elements of the sample spectral

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density, evaluated at a number of equidistant points, on certain trigo-
nometric functions using the method of weighted least squares. The
estimate seems to be different from the estimates derived in this paper.

The methods used in this paper are the Newton-Raphson and Scoring
Methods, applied to the maximum likelihood equations in the time and
frequency domains. The likelihood function is derived under the
assumption of normality of the data.

To summarize, Section 2 describes the model and the parameters to
be estimated. Sections 3 and 4 deal with the estimation problem in the
time and frequency domains respectively. Finally, in Section 5 we
derive the limiting average information matrix and show that the esti-
mates proposed are consistent and have the desired limiting multivariate
normal distribution, i.e. they are asymptotically efficient.
2. The Model

We have the observations \( y_1, y_2, \ldots, y_T \), where

\[
(2.1) \quad y_t = \varepsilon_t + A_1 \varepsilon_{t-1} + \ldots + A_q \varepsilon_{t-q}.
\]

Assumption 1: The \( \varepsilon_t \)'s are i.i.d. with mean zero and unknown covariance matrix \( \Sigma \).

Assumption 2: The roots of the determinantal equation

\[
(2.2) \quad |I + A_1 z + A_2 z^2 + \ldots + A_q z^q| = 0
\]

lie outside the unit circle.

Note: Assumption 2 enables us to recover the coefficients \( A_1, A_2, \ldots, A_q \) from the autocovariance matrices uniquely. The latter are the parameters of interest which are

\[
\Sigma_{(0)} = \varepsilon(y_t' y_t), \\
\Sigma_{(s)} = \varepsilon(y_{t-t+s}' y_t), \quad s=1, 2, \ldots, q.
\]

For ease of differentiation of the log likelihood function we should vectorize these matrices where we use the notation

\[
\text{Vec } A = \begin{pmatrix} a_1 \\ \vdots \\ a_n \end{pmatrix},
\]

where

\[
A = (a_1, \ldots, a_n)
\]

and the \( a_i \)'s are column vectors. But we notice that \( \Sigma_{(0)} \) is symmetric and hence should be treated separately. So we vectorize the diagonal and subdiagonal elements of \( \Sigma_{(0)} \) separately. So the parameters are
\[ \theta_0^{(1)} = \text{dg}(\Sigma^{(0)}) = \begin{pmatrix} c_{11}^{(0)} \\ \vdots \\ c_{pp}^{(0)} \end{pmatrix}, \]

\[ \theta_0^{(2)} = \widetilde{\text{vec}} \Sigma^{(0)}, \]

where the operator \( \widetilde{\text{vec}} \) vectorizes any matrix that it is applied to, ignoring the diagonal and upper diagonal elements of that matrix, e.g.,

\[ \widetilde{\text{vec}} \begin{pmatrix} 1 & 2 & 9 \\ 4 & 3 & 5 \\ 0 & 6 & 7 \end{pmatrix} = \begin{pmatrix} 4 \\ 0 \\ 6 \end{pmatrix}, \]

and finally

\[ \theta_s = \widetilde{\text{vec}} \Sigma^{(s)}, s=1, 2, \ldots, q. \]

Now we can put all these vectors in a single \( [qp^2 + \frac{p(p+1)}{2}] \times 1 \) vector \( \tilde{\theta} \), i.e.

\[ \tilde{\theta}' = (\theta_0', \theta_1', \ldots, \theta_q') \]

where

\[ \theta_0' = (\theta_0^{(1)}, \theta_0^{(2)}). \]

**Remarks.**

(i) We can find a matrix \( B \) such that

\[ B^{p \times p^2} \]

\[ \text{dg}(A) = B \widetilde{\text{vec}} A. \]

(2.3)
$B$ is obtained from the $p^2 \times p^2$ identity matrix by deleting all the rows except 1st, $p+2$nd, $2p+3$rd, ..., $p^2$th, i.e.

\[
B = \begin{pmatrix}
e'_1 \\
e'_p+2 \\
e'_2p+3 \\
\vdots \\
e'_2p
\end{pmatrix},
\]

(2.4)

where

\[
\tilde{\mathbf{I}}_{\frac{p^2}{2}} = (e'_1, e'_2, \ldots, e'_p) .
\]

(ii) In a similar manner we can find a $\frac{p(p-1)}{2} \times p^2$ matrix $C$ such that

\[
\tilde{\text{Vec}} \ 	ilde{A} = C \text{ Vec} \ 	ilde{A} .
\]

(2.5)

$C$ is obtained from $\tilde{\mathbf{I}}_{\frac{p^2}{2}}$ by deleting the following rows

1, $p+1$, $2p+1$, ..., $(p-1)p+1$
$p+2$, $2p+2$, ..., $(p-1)p+2$
$2p+3$, ..., $(p-1)p+3$
\vdots
$(p-1)p+p$ .

We shall find it convenient to introduce another vector $\tilde{\bar{\theta}}$ where

\[
\tilde{\theta}' = (\text{Vec'} \ \tilde{\Sigma}^{(0)}, \text{Vec'} \ \tilde{\Sigma}^{(1)}, \ldots, \text{Vec'} \ \tilde{\Sigma}^{(q)})
\]

(2.6)

\[
= (\tilde{\bar{\theta}}'_0, \tilde{\bar{\theta}}'_1, \ldots, \tilde{\bar{\theta}}'_q) .
\]
3. Estimation in the Time Domain

3.1. Introduction

We are going to use maximum likelihood estimation and proceed as if the $\varepsilon_t$'s in Section 2 are normally distributed and we shall show later that the resulting estimates have the same limiting covariance matrix irrespective of $\varepsilon_t$'s being normal.

Let

$$y' = (y'_{z1}, \ldots, y'_{zT}) .$$

Then

$$y \sim N(0, \Sigma) ,$$

where

$$\Sigma = \begin{pmatrix}
\Sigma(0) & \Sigma(1) & \cdots & \Sigma(q) & 0 & 0 & \cdots & 0 \\
\Sigma'(1) & \Sigma(0) & \Sigma(1) & \cdots & \Sigma(q) & 0 & \cdots & 0 \\
\vdots & \vdots & \ddots & \ddots & \ddots & \ddots & \ddots & \vdots \\
\Sigma'(q) & \vdots & \ddots & \ddots & \ddots & \ddots & \ddots & \Sigma(q) \\
0 & \vdots & \ddots & \ddots & \ddots & \ddots & \ddots & \vdots \\
\vdots & \vdots & \ddots & \ddots & \ddots & \ddots & \ddots & \vdots \\
0 & 0 & \cdots & \Sigma'(q) & \cdots & \Sigma'(1) & \Sigma(0) & \Sigma(1)
\end{pmatrix} .$$
Now using Kronecker products and the $T \times T$ matrix

$$L = \begin{pmatrix} 0 & I_T \otimes L^{-1} \\ I_T & 0 \end{pmatrix}$$

similar to the one introduced by Anderson (1975) we can write $\Sigma$ as

$$(3.1.1) \quad \Sigma = L_T \otimes \Sigma_0 + (L_T \otimes \Sigma_1 + L_T \otimes \Sigma_1') + \ldots + (L_T \otimes \Sigma_q + L_T \otimes \Sigma_q') .$$

The log likelihood of $y$ is

$$\log l(y) = -\frac{T}{2} \log 2\pi - \frac{1}{2} \log |\Sigma| - \frac{1}{2} y'y^{-1} .$$

The maximum likelihood estimates are a set of the roots of the equation

$$\frac{\partial \log l(y)}{\partial \theta} = 0 .$$

So we proceed to find the first derivative of the log likelihood, and in doing so we use the fact that

$$T \log l(y) = -\frac{1}{2} \frac{\partial}{\partial \theta} y'y^{-1} + \frac{1}{2} \left( \frac{\partial}{\partial \theta} y'y^{-1} - y'y^{-1} \right) .$$

$$(3.1.2) \quad \frac{\partial \log l(y)}{\partial \theta} = -\frac{1}{2} \frac{\partial}{\partial \theta} y'y^{-1} + \frac{1}{2} \left( \frac{\partial}{\partial \theta} y'y^{-1} - y'y^{-1} \right) .$$

### 3.2. The First Derivative of $\log l(y)$

As noted above we only need to find $\frac{\partial}{\partial \theta} y'y^{-1}$. It is more convenient to find $\frac{\partial}{\partial \theta} y'y^{-1} y$ and then, noting that
(3.2.1) \[ \frac{\partial}{\partial \sigma_{ij}} \tilde{y}' \Sigma^{-1} \tilde{y} = -\tilde{y}' \Sigma^{-1} \frac{\partial \Sigma}{\partial \sigma_{ij}} \tilde{y} \]
\[= -\tilde{y}' \Sigma^{-1} (\Sigma_T \otimes \Sigma_{ij}) \Sigma^{-1} \tilde{y} \]
\[= -\tilde{y}' \Sigma^{-1} (\Sigma_T \otimes \Sigma_{ji}) \Sigma^{-1} \tilde{y} \]
\[= \frac{\partial}{\partial \sigma_{ij}} \tilde{y}' \Sigma^{-1} \tilde{y}, \]

where \( \Sigma_{ij} \) is a matrix with 1 for the \( ij \)th element and 0's elsewhere, we get

(3.2.2) \[ \frac{\partial}{\partial \tilde{\sigma}_{ij}} \tilde{y}' \Sigma^{-1} \tilde{y} = 2 \frac{\partial}{\partial \sigma_{ij}} \tilde{y}' \Sigma^{-1} \tilde{y}, \]

where \( \frac{\partial}{\partial \tilde{\sigma}_{ij}} \) indicates that we take the symmetry of \( \Sigma^{(0)} \) into account.

The end result is that

(3.2.3) \[ \frac{\partial}{\partial \tilde{\theta}} \tilde{y}' \Sigma^{-1} \tilde{y} = \mathcal{G} \frac{\partial}{\partial \tilde{\theta}} \tilde{y}' \Sigma^{-1} \tilde{y}, \]

where \( \mathcal{G} \) is a \([qp^2 + \frac{p(p+1)}{2}] \times (q+1)p^2 \) matrix which can be written as

(3.2.4) \[ \mathcal{G} = \begin{pmatrix} \mathcal{G}_1 \\ \mathcal{G}_2 \\ \mathcal{G}_3 \end{pmatrix} \begin{pmatrix} p \\ \frac{p(p-1)}{2} \\ qp^2 \end{pmatrix}, \]

and

\[ \mathcal{G}_1 = (B, 0), \quad \mathcal{G}_2 = (2G, 0), \quad \mathcal{G}_3 = (0, \mathbb{I}), \]

with \( B \) and \( G \) as in Section 2.
So we proceed to find $\frac{\partial}{\partial \theta} y' \Sigma^{-1} y$, using the fact

$$\frac{\partial \Sigma^{-1}}{\partial x} = -\Sigma^{-1} \frac{\partial \Sigma}{\partial x} \Sigma^{-1}.$$ 

We shall also find it convenient to express $y' \Sigma^{-1} y$ differently using the identity

$$\text{Vec}(ABC) = (C' \otimes A) \text{Vec } B,$$

[see Minc and Marcus (1964)] which enables us to write

$$y' \Sigma^{-1} y = \text{Vec}(y' \Sigma^{-1} y) = (y' \otimes y') \text{Vec } \Sigma^{-1}.$$ 

So we only have to differentiate $\text{Vec } \Sigma^{-1}$, but

$$\frac{\partial \text{Vec } \Sigma^{-1}}{\partial x} = -\text{Vec} \left( \Sigma^{-1} \frac{\partial \Sigma}{\partial x} \Sigma^{-1} \right) = -\left( \Sigma^{-1} \otimes \Sigma^{-1} \right) \text{Vec } \frac{\partial \Sigma}{\partial x}.$$ 

Now

$$\frac{\partial \Sigma}{\partial \theta}^{(0)}_{ij} = I_T \otimes E_{ij}, \quad i, j=1, \ldots, p,$$

and

$$\frac{\partial \Sigma}{\partial \theta}^{(s)}_{ij} = L_s \otimes E_{ij} + L_s' \otimes E_{ji}, \quad i, j=1, \ldots, p.$$ 

So

$$\frac{\partial \text{Vec } \Sigma^{-1}}{\partial \theta}^{(0)}_{ij} = -\left( \Sigma^{-1} \otimes \Sigma^{-1} \right) \text{Vec} (L_T \otimes E_{ij}), \quad i, j=1, \ldots, p$$

$$= -\left( \Sigma^{-1} \otimes \Sigma^{-1} \right) \rho^{(0)}_{ij} \text{ say},$$
which yields

\[ \frac{\partial \text{Vec } \Sigma^{-1}}{\partial \theta'} \bigg|_{\tilde{\theta}_0} = -(\Sigma^{-1} \otimes \Sigma^{-1}) E_0 , \]

where

\[ E_0 = (\rho_{11}^{(0)}, \rho_{21}^{(0)}, \ldots, \rho_{p1}^{(0)}) . \]

Similarly, differentiating w.r.t. \( \sigma_{ij}^{(s)} \) we get

\[ \frac{\partial \text{Vec } \Sigma^{-1}}{\partial \sigma_{ij}^{(s)}} = -(\Sigma^{-1} \otimes \Sigma^{-1}) \text{Vec}(L^s \otimes E_{ij} + L'^s \otimes E_{j1}) \]

\[ = -(\Sigma^{-1} \otimes \Sigma^{-1}) \rho_{ij}^{(s)} , \quad i, j = 1, \ldots, p , \]

which yields

\[ \frac{\partial \text{Vec } \Sigma^{-1}}{\partial \tilde{\theta}_s} = -(\Sigma^{-1} \otimes \Sigma^{-1}) E_s , \quad s = 1, \ldots, q , \]

where

\[ E_s = (\rho_{11}^{(s)}, \rho_{21}^{(s)}, \ldots, \rho_{pp}^{(s)}) . \]

From (3.2.5) and (3.2.7) we get

\[ \frac{\partial \text{Vec } \Sigma^{-1}}{\partial \tilde{\theta}'} = -(\Sigma^{-1} \otimes \Sigma^{-1}) E , \]

where

\[ E = (E_0, E_1, \ldots, E_q) , \]

which in turn yields
(3.2.10) \[ \frac{\partial y' \Sigma^{-1} y}{\partial \theta'} = -(y'y^{-1} \otimes y'y^{-1}) \Sigma \).

Now to complete the computation of \[ \frac{\partial \log \ell(y)}{\partial \theta} \] we have to take the expectation of (3.2.10). We note that for any two vectors \( y, y' \) (of the same dimension) we have

\[ \text{Vec}(uv') = v \otimes u. \]

This means

\[ (y' \Sigma^{-1} \otimes y'y^{-1}) = \text{Vec}'(\Sigma^{-1}yy'y^{-1}). \]

So

\[ \mathcal{E}(y' \Sigma^{-1} \otimes y'y^{-1}) = \text{Vec}'(\Sigma^{-1}yy'y^{-1}) \]

\[ = \text{Vec}' \Sigma^{-1}. \]

Therefore

\[ \frac{\partial \log \ell(y)}{\partial \theta} = \frac{1}{2} (y' \Sigma^{-1} \otimes y'y^{-1} - \text{Vec}' \Sigma^{-1}) \Sigma, \]

or

\[ \frac{\partial \log \ell(y)}{\partial \theta} = \frac{1}{2} \mathcal{E}'(\Sigma^{-1}y \otimes \Sigma^{-1}y - \text{Vec} \Sigma^{-1}). \]

Finally

(3.2.11) \[ \frac{\partial \log \ell(y)}{\partial \theta} = \frac{1}{2} \mathcal{E}'(\Sigma^{-1}y \otimes \Sigma^{-1}y - \text{Vec} \Sigma^{-1}). \]

Note: We have from above

\[ \frac{\partial \log \ell(y)}{\partial \tilde{\theta}} = \frac{1}{2} \mathcal{E}'(\Sigma^{-1}y \otimes \Sigma^{-1}y - \text{Vec} \Sigma^{-1}). \]
Now

\[ E'(Y^{-1} \otimes \tilde{Y}^{-1}) = \begin{pmatrix} y'Y^{-1}A_1 \tilde{Y}^{-1}Y \\ \vdots \\ y'Y^{-1}A_{ip} \tilde{Y}^{-1}Y \end{pmatrix}, \]

where Vec \( A_{ir} \) is the \( r \)th column of \( E_i \) and \( A_{ir} \) is \( T_p \times T_p \). So we get

\[ E'(Y^{-1} \otimes \tilde{Y}^{-1}) = \left( I_p \otimes y'Y^{-1} \right) A_{1p}^{-1} \tilde{Y}^{-1}y, \]

where

\[ A_{1} = (A'_{11}, \ldots, A'_{ip}) \).

This means

\[ \frac{\partial \log Z(y)}{\partial \tilde{\theta}_i} = \frac{1}{2} \left( I_p \otimes y'Y^{-1} \right) A_{1p}^{-1} \tilde{Y}^{-1}y - \frac{1}{2} E'_i \text{Vec} \tilde{Y}^{-1}. \]

We shall use this latter form for \( \frac{\partial \log Z(y)}{\partial \tilde{\theta}_i} \) when finding the second derivative of \( \log Z(y) \).

3.3. The Numerical Approximations

The equation

\[ \frac{\partial \log Z(y)}{\partial \tilde{\theta}} = 0 \]

is clearly nonlinear and cannot be solved explicitly. So we will use numerical approximations to get asymptotically efficient estimates.

The methods we are going to use are the Newton-Raphson method and the
Scoring method. In both methods we need an initial consistent estimate (of order $T^2$) of $\hat{\theta}_0$, call it $\tilde{\hat{\theta}}_0(0)$, then the Newton-Raphson method consists of solving the following system of linear equations for $\tilde{\hat{\theta}}_0(1)$,

\[
(3.3.1) \quad - \frac{\partial^2 \log l(y)}{\partial \tilde{\theta} \partial \tilde{\theta}'} \bigg|_{\tilde{\hat{\theta}}_0(0)} (\tilde{\hat{\theta}}_0(1) - \tilde{\hat{\theta}}_0(0)) = \frac{\partial \log l(y)}{\partial \tilde{\theta}} \bigg|_{\tilde{\hat{\theta}}_0(0)} .
\]

The Scoring method is the same as above with $\varepsilon \frac{\partial^2 \log l(y)}{\partial \tilde{\theta} \partial \tilde{\theta}'}$ replacing $\frac{\partial \log l(y)}{\partial \tilde{\theta}}$, i.e. we have the equations

\[
(3.3.2) \quad -\varepsilon \left( \frac{\partial^2 \log l(y)}{\partial \tilde{\theta} \partial \tilde{\theta}'} \right) \bigg|_{\tilde{\hat{\theta}}_0(0)} (\tilde{\hat{\theta}}_0(1) - \tilde{\hat{\theta}}_0(0)) = \frac{\partial \log l(y)}{\partial \tilde{\theta}} \bigg|_{\tilde{\hat{\theta}}_0(0)} .
\]

To find $\tilde{\hat{\theta}}_0(0)$ we estimate the covariance matrices by their sample analogues, so

\[
(3.3.3) \quad \tilde{\hat{\theta}}_0(s) = \frac{1}{T-s} \sum_{t=1}^{T-s} y_t y_t' , \quad s = 0, 1, \ldots, q .
\]

3.4. The Scoring Method

To arrive at the linear equations for this method we need to find $\varepsilon \left( \frac{\partial^2 \log l(y)}{\partial \tilde{\theta} \partial \tilde{\theta}'} \right)$, but we know that

\[
\varepsilon \left( \frac{\partial^2 \log l(y)}{\partial \tilde{\theta} \partial \tilde{\theta}'} \right) = \varepsilon \left( \frac{\partial \log l(y)}{\partial \tilde{\theta}} \cdot \frac{\partial \log l(y)}{\partial \tilde{\theta}'} \right)
\]

and from (3.2.11)

\[
(3.4.1) \quad \varepsilon \left( \frac{\partial \log l(y)}{\partial \tilde{\theta}} \cdot \frac{\partial \log l(y)}{\partial \tilde{\theta}'} \right) = \varepsilon \left( \Sigma^{-1} y \otimes \Sigma^{-1} y' - \text{Vec} \Sigma^{-1} (y' \Sigma^{-1} \otimes y' \Sigma^{-1} - \text{Vec}' \Sigma^{-1}) \right) \otimes \Sigma^{-1} .
\]
Now
\[ z \sim \Sigma^{-1} y \sim N(0, \Sigma^{-1}) \]
and we need
\[ \mathcal{E}(z \otimes z - \text{Vec} \Sigma^{-1})(z' \otimes z' - \text{Vec}' \Sigma^{-1}) \]
\[ = \mathcal{E}(zz' \otimes zz') - \text{Vec} \Sigma^{-1} \text{Vec}' \Sigma^{-1} \]
since \( \mathcal{E}(z \otimes z) = \text{Vec} \Sigma^{-1} \).

So, suppose \( u \sim N(0, D) \), we want to find \( \mathcal{E}(uu' \otimes uu') \). The \( i_j \)th block of \( uu' \otimes uu' \) is
\[ u_i u_j uu' \]

To find \( \mathcal{E}(u_i u_j uu') \) we use the result (Anderson (1958, p. 39)).
\[ \mathcal{E}(u_i u_j u_s u_s) = d_{ij} d_{rs} + d_{ij} d_{rs} + d_{is} d_{jr} \]
which yields
\[ \mathcal{E}(u_i u_j uu') = d_{ij} D + d_{ij} d'_{ij} + d_{ij} d'_{ij} \]
where
\[ D = (d_1, \ldots, d_n) \]
From this we can get
\[ \mathcal{E}(uu' \otimes uu') = D \otimes D + \begin{pmatrix} \vdots \\ d'_{-1} \\ \vdots \\ d'_{-n} \\ d'_{-1}d'_{-2} \end{pmatrix} \]

\[ + \begin{pmatrix} \vdots \\ d_{-1}d_{-2}d_{-3} \end{pmatrix} \]

Now, the last matrix was shown by Magnus and Neudecker (1977) to be equal to

\[ K_{-n} (D \otimes D) , \]

where

\[ K_{-n} = \begin{pmatrix} E_{-11}^{(n)} & E_{-21}^{(n)} & \cdots & E_{-n1}^{(n)} \\ \vdots & \vdots & \ddots & \vdots \\ E_{-1n}^{(n)} & E_{-2n}^{(n)} & \cdots & E_{-nn}^{(n)} \end{pmatrix} \]

and \( E_{-ij}^{(n)} \)'s are \( n \times n \) matrices as defined before. So

\[ \mathcal{E}(uu' \otimes uu') = (I_{\tilde{n}} + K_{\tilde{n}})(D \otimes D) + \text{Vec } D \text{ Vec}' D . \]

Now for our problem

\[ u = z, D = \Sigma^{-1}, \tilde{n} = Tp \]

which means

\[ (3.4.3) \ \mathcal{E}(zz' \otimes zz') = (I_{T_p^2} + K_{T_p})(\Sigma^{-1} \otimes \Sigma^{-1}) + \text{Vec } \Sigma^{-1} \text{ Vec}' \Sigma^{-1} . \]

Finally we get the average information matrix as \( 1/T \) times

\[ \mathcal{E}\left( \frac{\partial \log \mathcal{L}}{\partial \theta} , \frac{\partial \log \mathcal{L}}{\partial \theta'} \right) = T \mathcal{E}'(I_{T_p^2} + K_{T_p})(\Sigma^{-1} \otimes \Sigma^{-1}) \mathcal{E}' . \]
So the linear equations for the scoring method are

$$
(3.4.5) \quad \mathbf{G} \mathbf{E}'(\mathbf{I}_{T} \otimes \hat{\Sigma}^{-1}) \mathbf{E} \mathbf{G}' \mathbf{E} (\hat{\Theta}_{0} - \hat{\Theta}_{0}) = 2 \mathbf{G} \mathbf{E}'(\hat{\Sigma}^{-1}) \mathbf{y} \mathbf{E} \mathbf{G}'(\hat{\Sigma}^{-1}) \mathbf{y} - \text{Vec} \ (\hat{\Sigma}^{-1}).
$$

Once we get $\hat{\Theta}_{1}$, we could use that as $\hat{\Theta}_{0}$ in (3.4.5) and get a second iterate $\hat{\Theta}_{2}$, but for large samples this is not necessary.

3.5. The Newton-Raphson Method

3.5.1. Preliminaries

To write down the linear equations for this method we need the second derivative of the log likelihood function. To derive the latter we first derive $\frac{\partial^2}{\partial \hat{\Theta} \partial \hat{\Theta}'} \log \ell(y)$, using the form (3.2.12) for $\frac{\partial}{\partial \hat{\Theta}}$ and then use

$$(3.5.1.1) \quad \frac{\partial^2}{\partial \hat{\Theta} \partial \hat{\Theta}'} \log \ell = \mathbf{g} \frac{\partial^2}{\partial \hat{\Theta} \partial \hat{\Theta}'} \log \ell \mathbf{g}'.
$$

3.5.2. The Derivation of $\frac{\partial^2}{\partial \hat{\Theta} \partial \hat{\Theta}'} \log \ell$.

As in (3.2.12) we have

$$
\frac{\partial}{\partial \hat{\Theta}} \log \ell = \frac{1}{2} \left( \mathbf{I}_{T} \otimes \hat{\Sigma}^{-1} \right) \hat{A}_{I} \hat{\Sigma}^{-1} y - \frac{1}{2} \hat{E}_{i} \text{Vec} \hat{\Sigma}^{-1},
$$

where $i = 0, 1, \ldots, q$.

We have the derivative of the second term w.r.t. $\hat{\Theta}$ as

$$(3.5.2.1) \quad \frac{1}{2} \hat{E}_{i} \left( \hat{\Sigma}^{-1} \otimes \hat{\Sigma}^{-1} \right) \hat{E},$$

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using (3.2.8). It remains to find the derivative of the first term.

To do this we shall find the derivative w.r.t. \( \partial_{\tilde{T}_j} \) for \( i \geq j \) and using symmetry will complete the derivation. So we let

\[
\tau_1 \equiv \tau_1(\tilde{\theta}) = \frac{1}{2}(I_{\sim p} \otimes \Sigma^{-1}_{\sim \Sigma} \otimes A_i \Sigma^{-1}_{\sim \Sigma} y, \ i = 0, 1, \ldots, q
\]

and

\[
\tau' = (\tau'_0, \ldots, \tau'_q).
\]

Now

\[
\frac{\partial \tau_1}{\partial \sigma^{(0)}} = -\frac{1}{2}(I_{\sim p} \otimes \Sigma^{-1}_{\sim \Sigma} (I_{\sim T} \otimes E_{\sim uv}) \Sigma^{-1}_{\sim \Sigma} A_i \Sigma^{-1}_{\sim \Sigma} y
\]

\[
= -\frac{1}{2}(I_{\sim p} \otimes \Sigma^{-1}_{\sim \Sigma} A_i \Sigma^{-1}_{\sim \Sigma} (I_{\sim T} \otimes E_{\sim uv}) \Sigma^{-1}_{\sim \Sigma} y.
\]

Now using \((A \otimes B) (C \otimes D) = (AC \otimes BD)\), we can factor out \((I_{\sim p} \otimes \Sigma^{-1}_{\sim \Sigma})\) to the left and also can factor out \(\Sigma^{-1}_{\sim \Sigma} y\) to the right, which results in

\[
\frac{\partial \tau_1}{\partial \sigma^{(0)}} = -\frac{1}{2}(I_{\sim p} \otimes \Sigma^{-1}_{\sim \Sigma} \) \(C_{i0,uv} \Sigma^{-1}_{\sim \Sigma} y,
\]

where

\[
C_{i0,uv} = [I_{\sim p} \otimes (I_{\sim T} \otimes E_{\sim uv}) \Sigma^{-1}_{\sim \Sigma}] A_i + A_i \Sigma^{-1}_{\sim \Sigma} (I_{\sim T} \otimes E_{\sim uv}.
\]

This leads to

\[
\frac{\partial \tau_1}{\partial \sigma^{(0)}} = \left(\frac{\partial \tau_1}{\partial \sigma^{(0)}}, \frac{\partial \tau_1}{\partial \sigma^{(0)}}, \ldots, \frac{\partial \tau_1}{\partial \sigma^{(0)}}\right)
\]

\[
= -\frac{1}{2}(I_{\sim p} \otimes \Sigma^{-1}_{\sim \Sigma}) (C_{i0,11} \Sigma^{-1}_{\sim \Sigma} y, \ldots, C_{i0,pp} \Sigma^{-1}_{\sim \Sigma} y)
\]

\[
= -\frac{1}{2}(I_{\sim p} \otimes \Sigma^{-1}_{\sim \Sigma}) C_{i0} (I_{\sim p} \otimes \Sigma^{-1}_{\sim \Sigma} y,
\]

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where

\[ z_{i0} = (z_{i0,11}, z_{i0,21}, \ldots, z_{i0,pp}), \quad i = 0, 1, \ldots, q. \]

Similarly

\[ \frac{\partial z_{i1}}{\partial \tilde{\theta}_j} = -\frac{1}{2} \left( I_{2} \otimes y' \Sigma^{-1} \right) c_{ij} \left( I_{2} \otimes \Sigma^{-1} y \right), \]

where

\[ c_{ij} = (c_{ij,11}, c_{ij,21}, \ldots, c_{ij,pp}) \]

and

\[ z_{ij,uv} = \left[ I_{2} \otimes \left( L^{j}_{\sim} \otimes E_{uv} + L^{j}_{\sim} \otimes E_{vu} \right) \Sigma^{-1} \right] A_{i} + A_{i} \Sigma^{-1} \left( L^{j}_{\sim} \otimes E_{uv} + L^{j}_{\sim} \otimes E_{vu} \right), \]

\[ j \leq i, \quad i = 0, 1, \ldots, q, \]

\[ u, v = 1, 2, \ldots p. \]

Finally for \( j \geq i \) using symmetry we have

\[ \frac{\partial z_{ij}}{\partial \tilde{\theta}_j} = -\frac{1}{2} \left( I_{2} \otimes y' \Sigma^{-1} \right) c'_{ij} \left( I_{2} \otimes \Sigma^{-1} y \right). \]

Now, for \( j \geq i \) define

\[ z_{ij} = c'_{ij} \]

and define

\[ z = (z_{ij}), \quad i, j = 0, 1, \ldots, q, \]

then from above we have
\[ (3.5.2.2) \frac{\partial \tau}{\partial \bar{\theta}^j} = -\frac{1}{2} \left( \left( I_{p} \otimes y^i \Sigma^{-1} \right) C_{ij} \left( I_{p} \otimes \Sigma^{-1} y \right) \right) i, j = 0, 1, \ldots, q \]

\[ = -\frac{1}{2} \left[ I_{q+1} \otimes \left( I_{p} \otimes y^i \Sigma^{-1} \right) \right] C \left[ I_{q+1} \otimes \left( I_{p} \otimes \Sigma^{-1} y \right) \right]. \]

Now notice that

\[ I_{m} \otimes (I_{m} \otimes A) \equiv I_{mn} \otimes A. \]

So (3.5.2.2) becomes

\[ \frac{\partial \tau}{\partial \bar{\theta}^j} = -\frac{1}{2} \left( I_{p} \otimes \Sigma^{-1} \right) C \left( I_{p} \otimes \Sigma^{-1} y \right). \]

Finally from (3.5.2.1) we have

\[ (3.5.2.3) \frac{\partial^2 \log \ell}{\partial \bar{\theta}^j \partial \bar{\theta}^i} = -\frac{1}{2} \left( I_{p} \otimes \Sigma^{-1} \right) C \left( I_{p} \otimes \Sigma^{-1} y \right) \]

\[ + \frac{1}{2} \Sigma^{-1} \otimes \Sigma^{-1} E. \]

3.5.3. The Equations

From (3.5.2.3) using (3.5.1.1) we have

\[ \frac{\partial^2 \log \ell}{\partial \bar{\theta}^j \partial \bar{\theta}^i} = -\frac{1}{2} \left( I_{p} \otimes \Sigma^{-1} \right) C \left( I_{p} \otimes \Sigma^{-1} y \right) \]

\[ + \frac{1}{2} \Sigma^{-1} \otimes \Sigma^{-1} E. \]

So the equations for the Newton-Raphson method are

\[ \left\{ \left( I_{p} \otimes \Sigma^{-1} (0) \right) \hat{\Sigma} (0) \right\} \hat{\Sigma}^{-1} (0) \]

\[ \times \left( \hat{\theta} (1) - \hat{\theta} (0) \right) = \Sigma^{-1} \otimes \Sigma^{-1} y - \text{Vec} \hat{\Sigma}^{-1} (0). \]
3.6. Summary

In the preceding sections we have derived the linear equations for the estimates, using the Newton-Rapnson and Scoring methods. The two resulting estimates are asymptotically equivalent, but at this point it is not clear which one is easier to compute. To write down the equations we first have to invert the \( T_p \times T_p \) block Toeplitz matrix \( \hat{\Sigma}(0) \) and once we have the equations, we need to find the best method (computationally easiest) to solve them. These problems will be considered in future work.

To compare the computational problems with that of Reinsel (1976), we see that in the latter a matrix \( \hat{\Sigma} \) which is essentially of the same form and size as \( \hat{\Sigma}(0) \) has to be inverted \( \hat{\Sigma} = \sum_{i=0}^{q} A_i \otimes L_i \). But once \( \hat{\Sigma}^{-1} \) is computed, Reinsel has shown that the solutions to his equations are the same as some generalized least square estimates.
4. Estimation in the Frequency Domain

4.1. Introduction

For a stationary process \( \{z_t, t = 0, \pm 1, \ldots \} \) with covariances \( \gamma_s = \mathbb{E}(z_{t+s} z_t) \), \( s = 0, \pm 1, \ldots \), the spectral density matrix \( \gamma \) is defined as

\[
\gamma(\lambda) = \frac{1}{2\pi} \sum_{s=-\infty}^{\infty} \gamma_s e^{-i\lambda s}.
\]

The covariances can be recovered from \( \gamma(\lambda) \) via

\[
\gamma_s = \int_{-\pi}^{\pi} \gamma(\lambda) e^{is\lambda} d\lambda.
\]

The sample analogue of the spectral density, the periodogram, is defined as

\[
\hat{\gamma}(\lambda) = \frac{1}{2\pi} \sum_{t=-(T-1)}^{T-1} \hat{\gamma}_s e^{-i\lambda s},
\]

where \( \hat{\gamma}_s \) is the sample analogue of \( \gamma_s \), more precisely

\[
\hat{\gamma}_s = \frac{1}{T} \sum_{t=1}^{T-s} y_t y_{t+s}.
\]

We can also represent \( \hat{\gamma}(\lambda) \) in terms of the discrete Fourier transforms

\[
\hat{\gamma}(\lambda) = \hat{w}(\lambda) \hat{w}^*(\lambda),
\]

where

\[
\hat{w}(\lambda) = \frac{1}{\sqrt{2\pi T}} \sum_{n=1}^{T} z_n e^{in\lambda}.
\]

For a fuller treatment see Anderson (1971).

If the process \( \{z_t, t = 0, \pm 1, \ldots \} \) is Gaussian then the log likelihood function may be approximated by
\( (4.1.5) \quad -\frac{1}{2} \log |\Gamma| - \frac{1}{2} \sum_{t} \text{tr} [f^{-1}(\lambda_{t}) I(\lambda_{t})] \),

where

\[ \Gamma = \text{Cov}(z_{1}, \ldots, z_{T}), \]

\[ \lambda_{t} = \frac{2\pi t}{T}, \quad t = 0, 1, \ldots, T-1. \]

Whittle (1953, 1961) suggested this for the case \( p=1 \) and Dunsmuir and Hannan (1976) showed that this leads to efficient estimates for general \( p \).

For our problem \( \gamma \equiv z, D_{s} \equiv \Sigma^{(s)}, \Gamma = \Sigma, \) and we only have a finite number of nonzero covariances so

\( (4.1.6) \quad f(\lambda) = \frac{1}{2\pi} \sum_{q=-q}^{q} \Sigma(s) e^{-i\lambda s} \)

and

\[ I(\lambda) = \frac{1}{2\pi} \sum_{-T}^{T-1} \hat{D}_{s} e^{-i\lambda s}. \]

So the log likelihood can be approximated by

\( (4.1.7) \quad \log \ell \approx -\frac{1}{2} \log |\Sigma| - \frac{1}{2} \sum_{t} \text{tr} [f_{t}^{-1} I_{t}] \),

where \( f_{t} \equiv f(\lambda_{t}), I_{t} = I(\lambda_{t}). \) We shall use the same approximation methods as in the time domain and will also use

\( (4.1.8) \quad \frac{\partial \log \ell}{\partial \Theta} = \frac{3}{\Sigma} \frac{\partial \log \ell}{\partial \Theta}, \)

\( (4.1.9) \quad \epsilon \left( \frac{\partial \log \ell}{\partial \Theta} \right) = 0. \)
4.2. The Derivation of \( \frac{\partial \log \ell}{\partial \tilde{\theta}} \)

Using (4.1.7) and (4.1.9) we get

\[
(4.2.1) \quad \frac{\partial \log \ell}{\partial \tilde{\theta}} = -\frac{1}{t} \sum_{i=1}^{\infty} \left[ \frac{\partial}{\partial \tilde{\theta}} \text{tr}(\tilde{f}^{-1} \tilde{T}_i) - e \frac{\partial}{\partial \tilde{\theta}} \text{tr}(\tilde{f}^{-1} \tilde{T}_i) \right].
\]

Now

\[
\frac{\partial}{\partial \tilde{\theta}} \text{tr}(\tilde{f}^{-1} \tilde{T}_i) = -\text{tr}(\tilde{f}^{-1} \frac{\partial \tilde{f}}{\partial \tilde{x}} \tilde{f}^{-1} \tilde{T}_i).
\]

Differentiating (4.1.6) yields

\[
\frac{\partial \tilde{f}}{\partial \tilde{\sigma}^{(0)}} = \frac{1}{2\pi} E_{uv},
\]

\[
\frac{\partial \tilde{f}}{\partial \tilde{\sigma}^{(s)}} = \frac{1}{2\pi} \left[ e^{ts} E_{uv} + e^{ts} E_{vu} \right], \quad u,v = 1, \ldots, p.
\]

So

\[
\frac{\partial}{\partial \tilde{\sigma}^{(0)}} \text{tr}(\tilde{f}^{-1} \tilde{T}_i) = -\frac{1}{2\pi} \text{tr}(\tilde{f}^{-1} E_{uv} \tilde{f}^{-1} \tilde{T}_i)
\]

\[
= -\frac{1}{2\pi} \text{tr}(\tilde{f}^{-1} e^{e'} \tilde{f}^{-1} \tilde{T}_i)
\]

\[
= -\frac{1}{2\pi} e^{e'} \tilde{f}^{-1} \tilde{T}_i \tilde{f}^{-1} e^{e'}
\]

\[
= -\frac{1}{2\pi} \left( h_{vu} \right)_{uv}, \quad u,v = 1, \ldots, p,
\]

where

\[
(4.2.2) \quad h_{vu} = \tilde{f}^{-1} \tilde{T}_i \tilde{f}^{-1}.
\]
and we have used the fact that \( E_{uv} = e_u^t e_v^t \), where \( e_u^t, e_v^t \) are the \( u \th \) and \( v \th \) column of the \( p \times p \) identity matrix. Now we can easily see

\[
(4.2.3) \quad \frac{\partial}{\partial \theta^0_{-t}} \text{tr}(f^{-1}_{-t} I_{-t}) = -\frac{1}{2\pi} \text{Vec}(h_t^t).
\]

Similarly

\[
\frac{\partial}{\partial \theta_{uv}^{(s)}} \text{tr}(f^{-1}_{-t} I_{-t}) = -\frac{1}{2\pi} \text{tr} \left[ f^{-1}_{-t} (e^{i\lambda^s_t} E_{-uv} + e^{-i\lambda^s_t} E_{-vu}) f^{-1}_{-t} I_{-t} \right]
\]

\[
= -\frac{1}{2\pi} \left[ e^{i\lambda^s_t} h_t^t + e^{-i\lambda^s_t} h_t^t \right]_{uv}, \quad u,v = 1, \ldots, p,
\]

which yields

\[
(4.2.4) \quad \frac{\partial}{\partial \theta^s_{-t}} \text{tr}(f^{-1}_{-t} I_{-t}) = -\frac{1}{2\pi} \text{Vec}(e^{i\lambda^s_t} h_t^t + e^{-i\lambda^s_t} h_t^t), \quad s = 1, \ldots, q.
\]

To complete the derivation of \( \frac{\partial \log L}{\partial \theta} \) we have to take the expectation of (4.2.3) and (4.2.4). This yields

\[
\mathbb{E}\left[ \frac{\partial}{\partial \theta^0_{-t}} \text{tr}(f^{-1}_{-t} I_{-t}) \right] = -\frac{1}{2\pi} \text{Vec}(f^{-1}_{-t}),
\]

\[
\mathbb{E}\left[ \frac{\partial}{\partial \theta^s_{-t}} \text{tr}(f^{-1}_{-t} I_{-t}) \right] = -\frac{1}{2\pi} \text{Vec}\left[ e^{i\lambda^s_t} f^{-1}_{-t} + e^{-i\lambda^s_t} f^{-1}_{-t} \right],
\]

since \( \mathbb{E} h_t = f^{-1}_{-t} \), which follows from \( \mathbb{E}(I_{-t}) = f_{-t} + O(T^{-1}) \). Now let

\[
\ell_t = h_t - f^{-1}_{-t},
\]

then using (4.2.1) we have
\[ \frac{\partial \log \mathcal{L}}{\partial \theta_0} = \frac{1}{4\pi} \sum_t \text{Vec} \left[ \frac{i\lambda s}{t} \frac{\mathcal{L}}{s} + e^{i\lambda s} \frac{\mathcal{L}}{s} \right], \quad s = 1, \ldots, q. \]

Now, from (4.1.6) it is obvious that \( f' = \tilde{f} \) or \( f = \tilde{f}^* \), where \( * \) indicates "conjugate transpose", i.e. \( \tilde{f} \) is Hermitian. Also \( I'_z = \tilde{I}_z \), which leads to \( h' = \tilde{h} \) and \( \lambda' = \tilde{\lambda} \). We can use this to simplify (4.2.6) as follows:

\[ \frac{\partial \log \mathcal{L}}{\partial \theta} = \frac{1}{4\pi} \sum_t \text{Vec}(e^{i\lambda s} \mathcal{L}) + \frac{1}{4\pi} \sum_t \text{Vec}(e^{i\lambda s} \mathcal{L}) \]

\[ = \frac{1}{2\pi} \sum_t \text{Vec}(e^{i\lambda s} \mathcal{L}), \]

because the first sum is real. The reason for this is that

\[ e^{i\lambda t} = e^{-i\lambda t} = e^{-i\frac{2\pi t}{T}} = e^{i2\pi \frac{(T-t)}{T}} = e^{i\lambda(T-t)}. \]

This means that for any real function \( \eta(*) \)

\[ \sum_{t=0}^{T-1} \eta(e^{i\lambda t}) = \eta(0) + \sum_{t=0}^{\frac{T-1}{2}} \eta(e^{i\lambda t}) + \eta(e^{i\lambda t}) , \]

for \( T \) odd, and

\[ \sum_{t=0}^{T-1} \eta(e^{i\lambda t}) = \eta(0) + \sum_{t=0}^{\frac{T-1}{2}} \eta(e^{i\lambda t}) + \eta(e^{i\lambda t}) + \eta(-1) , \]

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for $T$ even. So $\sum_{t=0}^{T-1} \eta(e^{i\lambda_t})$ is real. The same argument allows us to rewrite (4.2.5) as

$$\frac{\partial \log \mathcal{L}}{\partial \tilde{\theta}_0} = \frac{1}{4\pi} \sum_t \text{Vec } \tilde{\mathcal{L}}_t.$$ 

Finally we can conclude

$$(4.2.7) \quad \frac{\partial \log \mathcal{L}}{\partial \tilde{\theta}_0} = \frac{1}{2\pi} \sum_t \text{J}_t \text{Vec } \tilde{\mathcal{L}}_t,$$

where

$$\text{J}_t = \left( I_2, e^{i\lambda_t} I_2, \ldots, e^{q_i\lambda_t} I_2 \right).$$

There is an alternative form for (4.2.7) which will be more useful in deriving the second derivative of $\log \mathcal{L}$. It is obtained by noticing that we can find a matrix $M$, such that

$$\text{Vec } \tilde{\mathcal{L}}_t = M \text{Vec } \tilde{\mathcal{L}}'_t,$$

which then enables us to rewrite (4.2.6) as

$$(4.2.8) \quad \frac{\partial \log \mathcal{L}}{\partial \tilde{\theta}_s} = \frac{1}{4\pi} \sum_t (e^{i\lambda_t^s} I_2 + e^{-i\lambda_t^s} M) \text{Vec } \tilde{\mathcal{L}}_t.$$ 

It is easily verified that in fact

$$M = K_{p},$$

where $K_p$ was defined in (3.4.2). So the alternative form for the first derivative of $\log \mathcal{L}$ is

$$(4.2.9) \quad \frac{\partial \log \mathcal{L}}{\partial \tilde{\theta}_s} = \frac{1}{4\pi} \sum_t H_t \text{Vec } \tilde{\mathcal{L}}_t.$$
where

\[ H'_{\sim t} = H'(\lambda_t) = (I_{t \sim p}, e^{i \lambda_t}_{\sim t} I_{t \sim p} + e^{-i \lambda_t}_{\sim t} K_{\sim p}, \ldots, e^{q_i \lambda_t}_{\sim t} I_{t \sim p} + e^{-p_i \lambda_t}_{\sim p}). \]

4.3. The Second Derivative of log \( \ell \)

Using (4.2.9) we have

\[ \frac{\partial^2 \log \ell}{\partial \theta \partial \theta'} = \frac{1}{4\pi} \sum_{t} \frac{\partial \text{Vec} \ell_{\sim t}}{\partial \theta'}. \]

Now

\[ \ell_{\sim t} = f_{-1}^{-1} I_{-t} f_{-1}^{-1} I_{-t} - f_{-1}^{-1}, \]

which yields

\[ \frac{\partial \ell_{\sim t}}{\partial x} = f_{-1}^{-1} \frac{\partial f_{-1}}{\partial x} f_{-1}^{-1} I_{-t} f_{-1}^{-1} I_{-t} - f_{-1}^{-1} I_{-t} f_{-1}^{-1} \frac{\partial f_{-1}}{\partial x} f_{-1}^{-1} I_{-t} f_{-1}^{-1} I_{-t} + f_{-1}^{-1} \frac{\partial f_{-1}}{\partial x} f_{-1}^{-1}, \]

which in turn yields

\[ \frac{\partial \text{Vec} \ell_{\sim t}}{\partial x} = [f_{-1}^{-1} \otimes f_{-1}^{-1} - f_{-1}^{-1} \otimes h_{\sim t} - h_{\sim t} \otimes f_{-1}^{-1}] \text{Vec} \frac{\partial f_{-1}}{\partial x}, \]

say. Now for \( x = \sigma_{uv}^{(0)} \) we get

\[ \frac{\partial \text{Vec} \ell_{\sim t}}{\partial \sigma_{uv}^{(0)}} = \frac{1}{2\pi} \phi_t \, \epsilon_{uv}, \]

where

\[ \epsilon_{uv} = \text{Vec} \, E_{uv}, \ u, v = 1, \ldots, p. \]
So

\[(4.3.3) \quad \frac{\partial \text{Vec } l}{\partial \theta'_t} \bigg|_{\sim 0} = \frac{1}{2\pi} \phi_t \left( e_{11}, e_{21}, \ldots, e_{pp} \right)\]

\[= \frac{1}{2\pi} \phi_t \frac{I}{p} \frac{1}{2} = \frac{1}{2\pi} \phi_t .\]

Similarly

\[\frac{\partial \text{Vec } l}{\partial \theta_{uv}(s)} = \frac{1}{2\pi} \phi_t (e_{uv} - i\lambda_t e_{uv} + i\lambda_t e_{uv}) , \quad u, v = 1, \ldots, p .\]

Now

\[e_{uv} = \text{Vec } E_{uv} = \text{Vec } E'_{uv} = K \text{Vec } E_{uv} = K e_{uv} ,\]

which means

\[\frac{\partial \text{Vec } l}{\partial \theta_{uv}(s)} = \frac{1}{2\pi} \phi_t \left( e_{uv} - i\lambda_t e_{uv} + i\lambda_t e_{uv} \right) , \quad u, v = 1, \ldots, p , \quad s = 1, \ldots, q .\]

Finally

\[(4.3.4) \quad \frac{\partial \text{Vec } l}{\partial \theta'_t} \bigg|_{\sim s} = \frac{1}{2\pi} \phi_t \left( e_{11} + e_{21} + i\lambda_t K \right) , \quad s = 1, \ldots, q .\]

So

\[(4.3.5) \quad \frac{\partial \text{Vec } l}{\partial \theta'_t} \bigg|_{\sim} = \frac{1}{2\pi} \phi_t \left( I_2 + e_{11} + e_{21} + \cdots + e_{q1} + e_{q2} + \cdots \right)\]

\[= \frac{1}{2\pi} \phi_t H^* .\]
Now we get the second derivative from (4.3.1)

\[
\frac{\partial^2 \log \ell}{\partial \bar{\theta} \partial \bar{\theta}'} = \frac{1}{8\pi^2} \sum_t \bar{H}_t \phi_t H_t^* .
\]

(4.3.6)

4.4. The Newton-Raphson Method

As in the time domain we will use \( \hat{\lambda}(s) \)'s as initial estimates of the covariances, and so

\[
\hat{f}_{\sim t}(0) = \frac{1}{2\pi} \sum_{s=-q}^{q} e^{-i\lambda_s} \hat{f}(s) \frac{\hat{\lambda}(s)}{\sim (0)}
\]

will be the initial estimate of \( \hat{f}_{\sim t}(0) \). Accordingly we form \( \hat{f}_{\sim t}(0)^{-1} \) and

\[
\hat{\phi}_{\sim t}(0) = \frac{\hat{f}_{\sim t}(0)^{-1}}{\hat{f}_{\sim t}(0)} - \hat{f}_{\sim t}(0)^{-1} \hat{f}_{\sim t}(0)^{-1} \hat{f}_{\sim t}(0)^{-1} .
\]

Using (4.2.7) we have

\[
\frac{\partial \log \ell}{\partial \bar{\theta}} = G \frac{\partial \log \ell}{\partial \bar{\theta}'} = \frac{1}{4\pi} \sum_t GJ_t \text{Vec} \ell_t .
\]

(4.4.1)

Similarly from (4.3.6) we get

\[
\frac{\partial^2 \log \ell}{\partial \bar{\theta} \partial \bar{\theta}'} = \frac{1}{8\pi^2} \sum_t \bar{H}_t \phi_t H_t^* G' .
\]

(4.4.2)

So the linear equations for the Newton-Raphson method are
\[(4.4.3) \quad \left(- \sum_t \hat{\phi}_t (0) \frac{\partial}{\partial \hat{\phi}_t} (0) \mathbf{H}_t^* \mathbf{C}' \right) \left( \hat{\theta}_t (0) - \hat{\theta}_t (1) \right) = 2 \sum_t \mathbf{G}_t \mathbf{J} \mathbf{V} \mathbf{e}_t (0). \]

4.5. The Scoring Method

As mentioned earlier, to get the linear equations for this method we replace \( \frac{\partial^2 \log \mathcal{L}}{\partial \hat{\theta} \partial \hat{\theta}'} \) by its expectation in the Newton-Raphson method. The latter is

\[
\mathbb{E} \left( \frac{\partial^2 \log \mathcal{L}}{\partial \hat{\theta} \partial \hat{\theta}'} \right) = \frac{1}{8 \pi} \sum_t \mathbf{G}_t \mathbf{E} (\phi_t) \mathbf{H}_t^* \mathbf{C}'.
\]

Now

\[
\mathbb{E} (I_{\hat{\ell}_t}) \sim \mathbf{f}_t,
\]

which leads to

\[
\mathbb{E} (\mathbf{h}_t) = \mathbb{E} (\mathbf{f}^{-1}_t \mathbf{I}_t \mathbf{f}^{-1}_t) \sim \mathbf{f}^{-1}_t,
\]

which in turn leads to

\[
\mathbb{E} (\phi_t) \sim - (\mathbf{f}^{-1}_t \otimes \mathbf{f}^{-1}_t)
\]

So

\[
\mathbb{E} \left( \frac{\partial^2 \log \mathcal{L}}{\partial \hat{\theta} \partial \hat{\theta}'} \right) = - \frac{1}{8 \pi} \sum_t \mathbf{G}_t \mathbf{E} (\mathbf{f}^{-1}_t \mathbf{I}_t \mathbf{f}^{-1}_t \mathbf{H}_t^* \mathbf{C}')
\]

ignoring the terms of order \( T^{-1} \). Finally the linear equations for this method are

\[
\left[ \sum_t \mathbf{G}_t \mathbf{H}_t (\mathbf{f}^{-1}_t \otimes \mathbf{f}^{-1}_t \mathbf{H}_t^* \mathbf{C}') \mathbf{E} (\hat{\theta}_t (0) - \hat{\theta}_t (1) \mathbf{e}_t (0)) \right] = 2 \sum_t \mathbf{G}_t \mathbf{J} \mathbf{V} \mathbf{e}_t (0).
\]
4.5. Remarks

As in the time domain, there are computational problems to be considered in setting up and solving the equations derived in the preceding sections. In this case the main problem in setting up the equations is the inversion of \( p \times p \) matrices \( \hat{f}_t^{\ast}(0) \), \( t = 1, \ldots, T-1 \). And again we have to find the best way to solve the resulting equations. It seems that the computation of the estimates in the time domain is easier than that in the frequency domain.

Comparing with Nicholls (1976) and Anderson (1978) which deal with the estimation of the coefficients \( A_i \), \( i = 1, \ldots, q \), we see that the main problem of inversion of \( \hat{f}_t^{\ast}(0) \)'s is also present in these papers.
5. Asymptotic Properties

The four estimates proposed in this paper are asymptotically equivalent and we shall show that they are efficient, i.e.

\[ \sqrt{T} \left( \hat{\theta}_p - \theta \right) \Rightarrow N(0, J(\theta)) \cap L, \]

where \( J(\theta) \) is the limiting average information matrix and \( \Rightarrow \) indicates convergence in distribution.

To find \( J(\theta) \), by definition we have

\[
J(\theta) = \lim_{T \to \infty} -\frac{1}{T} \mathbb{E} \left( \frac{\partial^2 \log L}{\partial \theta \partial \theta} \right)_{\bar{\theta}} \]

\[
= \lim_{T \to \infty} \frac{1}{T} \left\{ \frac{1}{8\pi^2} \sum_{t} \mathcal{G} \left( f_1 \otimes f_1 \right)^*_{\bar{c}} \right\},
\]

\[
= \frac{1}{16\pi^3} \int_{0}^{2\pi} \mathcal{G} \left( f_1 \otimes f_1 \right)^*_{\bar{c}} \, d\lambda,
\]

where the argument \( \lambda \) is omitted from \( \mathcal{G} \) and \( f_1 \).

The four estimates are obtained from equations like

\[
\hat{J}(\theta) = \hat{J} \left( \hat{\theta}_p - \theta \right) \Rightarrow J(\theta),
\]

(5.2)

\[
\hat{J}(\theta) = \frac{1}{T} \frac{\partial \log L}{\partial \theta} \bigg|_{\theta = \hat{\theta}_p - \theta},
\]

where \( \hat{J}(\theta) \) is a consistent estimate of \( J(\theta) \). We can rewrite (5.2) as

\[
\hat{J}(\theta) = \hat{J} \left( \hat{\theta}_p - \theta \right) \Rightarrow J(\theta),
\]

(5.3)

\[
\hat{J} \left( \hat{\theta}_p - \theta \right) \Rightarrow J(\theta),
\]

where \( \theta \) is the true parameter. Now

\[
\frac{1}{\sqrt{T}} \frac{\partial \log L}{\partial \theta} = \frac{1}{\sqrt{T}} \frac{\partial \log L}{\partial \theta} \bigg|_{\theta = \hat{\theta}_p - \theta} + \frac{1}{\sqrt{T}} \frac{\partial^2 \log L}{\partial \theta^2} \bigg|_{\theta = \hat{\theta}_p - \theta} \left( \theta - \hat{\theta}_p - \theta \right),
\]

(5.4)
where $|\hat{\theta} - \hat{\theta}_0^+| < |\hat{\theta} - \hat{\theta}_0^-|$. Now (5.3) can be rewritten using (5.4)

\begin{equation}
\sqrt{T}(\hat{\theta}_{\sim(0)} - \theta) = \left[ \mathcal{J}(\theta_{\sim(0)}) + \frac{1}{T} \frac{\partial^2 \log \mathcal{L}}{\partial \bar{\theta} \partial \bar{\theta}} \right]_{\theta = \bar{\theta}}^+ \mathcal{J}(\theta_{\sim(0)}) + \frac{1}{\sqrt{T}} \frac{\partial \log \mathcal{L}}{\partial \bar{\theta}}.
\end{equation}

Now noticing that

\begin{equation}
\frac{1}{T} \frac{\partial^2 \log \mathcal{L}}{\partial \bar{\theta} \partial \bar{\theta}} \right|_{\theta = \bar{\theta}}^+ \mathcal{J}(\theta_{\sim(0)})
\end{equation}

and $\sqrt{T}(\hat{\theta}_{\sim(0)} - \theta)$ is bounded in probability, we see that (5.5) is (asymptotically) equivalent to

\begin{equation}
\sqrt{T}(\hat{\theta}_{\sim(1)} - \theta) = \mathcal{J}^{-1}(\theta_{\sim(0)}) \frac{1}{\sqrt{T}} \frac{\partial \log \mathcal{L}}{\partial \bar{\theta}}.
\end{equation}

**Theorem.** Under (2.1) and Assumptions 1 and 2 of Section 2,

\begin{equation}
\sqrt{T}(\hat{\theta}_{\sim(1)} - \theta) \Rightarrow N(0, \mathcal{J}^{-1}(\theta)),
\end{equation}

where $\hat{\theta}_{\sim(1)}$ is any one of the four estimates derived in this paper.

**Proof.** Using (5.6), it suffices to show

\begin{equation}
\frac{1}{\sqrt{T}} \frac{\partial \log \mathcal{L}}{\partial \bar{\theta}} \Rightarrow N(0, \ast).
\end{equation}

Let

\begin{equation}
\xi = \begin{pmatrix}
\text{Vec}(A_1, \ldots, A_q) \\
\text{dg \text{\text{\text{v}}}} \\
\text{Vec \text{\text{\text{v}}}}
\end{pmatrix},
\end{equation}

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where $A_i$'s and $V$ were defined in Section 2. Now

$$\frac{\partial \log L}{\partial \xi_1} = \sum_j \frac{\partial \log L}{\partial \theta_j} \cdot \frac{\partial \theta_j}{\partial \xi_1},$$

which means

$$\frac{\partial \log L}{\partial \xi} = \frac{\partial \theta'}{\partial \xi} \cdot \frac{\partial \log L}{\partial \theta}.$$

It follows from Assumption 2 of Section 2 that $\frac{\partial \theta'}{\partial \xi}$ is nonsingular, which means

$$(5.7) \quad \frac{1}{\sqrt{T}} \frac{\partial \log L}{\partial \theta} = \left(\frac{\partial \theta'}{\partial \xi}\right)^{-1} \frac{1}{\sqrt{T}} \frac{\partial \log L}{\partial \xi}.$$

But it has been shown by Nicholls (1976) and Reinsel (1976) that

$$\sqrt{T} (\xi - \bar{\xi}) \overset{d}{\rightarrow} N(0, \cdot),$$

which is the same as

$$\frac{1}{\sqrt{T}} \frac{\partial \log L}{\partial \xi} \overset{d}{\rightarrow} N(0, \cdot).$$

So (5.7) gives us

$$(5.8) \quad \frac{1}{\sqrt{T}} \frac{\partial \log L}{\partial \theta} \overset{d}{\rightarrow} N(0, \cdot).$$

The limiting covariance matrix in (5.8) is obviously $\widehat{J}(\theta)$, so

$$\sqrt{T} (\hat{\theta}(1) - \theta) \overset{d}{\rightarrow} N(0, \widehat{J}^{-1}(\theta)).$$

Q.E.D.
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**MAXIMUM LIKELIHOOD ESTIMATION OF THE COVARIANCES OF THE VECTOR MOVING AVERAGE MODELS IN THE TIME AND FREQUENCY DOMAINS**

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**KEY WORDS**
Maximum likelihood estimation, vector moving average models, Newton-Raphson and Scoring iterative procedures, Time and Frequency Domains.

**ABSTRACT**
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20. ABSTRACT.

The vector moving average process is a stationary stochastic process \( \{y_t\} \) satisfying \( y_t = \sum_{i=0}^{q} A_i \varepsilon_{t-i} \), where the unobservable process \( \{\varepsilon_t\} \) consists of independently identically distributed random variables. The matrix parameters \( \sum(s) = \sum_{t=t+s} y_t y_s' \), \( s = 0, 1, \ldots, q \) are estimated from the observations \( y_1, \ldots, y_T \). The likelihood function is derived under normality and to solve the maximum likelihood equations the Newton-Raphson and Scoring methods are used. The estimation problem is considered in the time and frequency domains. Asymptotic efficiency of the estimates is established.